
Counting and Locating Multiple Solutions of Estimating Equations

Speaker: Donald Richards (Penn State University)

This talk is based on joint work with:

Despina Stasi (Penn State University)

Elizabeth Gross (NC State University)

Sonja Petrović (Illinois Institute of Technology)

Logistic regression

θ_i : The probability that individual i in a random sample of n individuals will develop a particular characteristic during a follow-up period.

Y_i : Bernoulli random variable which indicates whether or not individual i develops the characteristic.

Y_1, \dots, Y_n are assumed independent, so they have joint p.d.f.

$$f(y_1, \dots, y_n; \theta_1, \dots, \theta_n) = \prod_{i=1}^n \theta_i^{y_i} (1 - \theta_i)^{1-y_i}, \quad y_i = 0 \text{ or } 1$$

List the individuals so that the first m are those who have the characteristic; so, $y_i = 1, i \leq m$, and $y_i = 0, i > m$.

Likelihood function:

$$L(\theta_1, \dots, \theta_n) = \prod_{i=1}^m \theta_i \cdot \prod_{i=m+1}^n (1 - \theta_i)$$

Predictor variables: x_1, x_2, \dots, x_k (and $x_0 \equiv 1$)

Data: x_{ij} , the observed value of x_j for the i th individual.

$\beta = (\beta_0, \beta_1, \dots, \beta_k)$: A vector of unknown parameters to be estimated by the method of maximum likelihood.

Model θ_i through a *logistic relationship*:

$$\theta_i = \frac{1}{1 + e^{-\sum_{j=0}^k \beta_j x_{ij}}}$$

The likelihood function:

$$L(\boldsymbol{\beta}) = \prod_{i=1}^m \frac{1}{1 + e^{-\sum_{j=0}^k \beta_j x_{ij}}} \cdot \prod_{i=m+1}^n \frac{1}{1 + e^{\sum_{j=0}^k \beta_j x_{ij}}}$$

The derivatives of $\log L(\boldsymbol{\beta})$ w.r.t. β_r , $r = 0, \dots, k$:

$$\frac{\partial}{\partial \beta_r} \log L(\boldsymbol{\beta}) = \sum_{i=1}^m x_{ir} \frac{e^{-\sum_{j=0}^k \beta_j x_{ij}}}{1 + e^{-\sum_{j=0}^k \beta_j x_{ij}}} - \sum_{i=m+1}^n x_{ir} \frac{e^{\sum_{j=0}^k \beta_j x_{ij}}}{1 + e^{\sum_{j=0}^k \beta_j x_{ij}}}$$

The system of $k + 1$ likelihood equations:

$$\sum_{i=1}^m \frac{1}{1 + e^{\sum_{j=0}^k \beta_j x_{ij}}} \begin{pmatrix} x_{i0} \\ x_{i1} \\ \vdots \\ x_{ik} \end{pmatrix} = \sum_{i=m+1}^n \frac{e^{\sum_{j=0}^k \beta_j x_{ij}}}{1 + e^{\sum_{j=0}^k \beta_j x_{ij}}} \begin{pmatrix} x_{i0} \\ x_{i1} \\ \vdots \\ x_{ik} \end{pmatrix}$$

Change of variables:

$$\gamma_j \equiv e^{\beta_j}, \quad j = 0, \dots, k$$

The likelihood equations: For $\gamma_0, \dots, \gamma_k > 0$,

$$\sum_{i=1}^m \frac{1}{1 + \gamma_0^{x_{i0}} \dots \gamma_k^{x_{ik}}} \begin{pmatrix} x_{i0} \\ x_{i1} \\ \vdots \\ x_{ik} \end{pmatrix} = \sum_{i=m+1}^n \frac{\gamma_0^{x_{i0}} \dots \gamma_k^{x_{ik}}}{1 + \gamma_0^{x_{i0}} \dots \gamma_k^{x_{ik}}} \begin{pmatrix} x_{i0} \\ x_{i1} \\ \vdots \\ x_{ik} \end{pmatrix}$$

Problems:

1. Count the number of solutions of this system of equations?
2. Can we calculate *all* solutions?

The Donner party data

Row 1: Age

Row 2: Sex (1=male, 0=female)

Survived vs. Died

40	40	28	22	23	28	15	20	18	25	20	32	32	24	30
0	1	1	0	0	1	0	0	1	1	1	1	0	0	1
21	46	32	23	25	23	30	28	40	45	62	65	45	25	28
0	1	0	1	0	1	1	1	1	0	1	1	0	0	1
23	47	57	25	60	15	50	25	30	25	25	25	30	35	24
1	0	1	1	1	1	0	1	1	1	1	1	1	1	1

Suppose we were given the data on individuals 8, 10, 29, and 43 only, then the system of likelihood equations is:

$$\begin{pmatrix} 1 & 1 & 1 & 1 \\ 20 & 25 & 25 & 30 \\ 0 & 1 & 0 & 1 \end{pmatrix} \begin{pmatrix} a \\ b \\ c \\ d \end{pmatrix} = 0,$$

where $\gamma_0, \gamma_1, \gamma_2 > 0$ and

$$a = \frac{1}{1 + \gamma_0 \gamma_1^{20} \gamma_2^0}, \quad b = \frac{1}{1 + \gamma_0 \gamma_1^{25} \gamma_2^1},$$
$$c = -\frac{\gamma_0 \gamma_1^{25} \gamma_2^0}{1 + \gamma_0 \gamma_1^{25} \gamma_2^0}, \quad d = -\frac{\gamma_0 \gamma_1^{30} \gamma_2^1}{1 + \gamma_0 \gamma_1^{30} \gamma_2^1}.$$

Row-reduction leads to: $a = -b = -c = d$, so $ab < 0, cd < 0$.

Conclusion: The likelihood equations have no real solutions.

Suppose we were given the data on individuals 2, 20, 24, and 29 only. Then the likelihood equations are

$$\begin{pmatrix} 1 & 1 & 1 & 1 \\ 40 & 25 & 40 & 25 \\ 1 & 0 & 1 & 0 \end{pmatrix} \begin{pmatrix} a \\ b \\ c \\ d \end{pmatrix} = 0$$

where $\gamma_0, \gamma_1, \gamma_2 > 0$ and

$$a = \frac{1}{1 + \gamma_0 \gamma_1^{40} \gamma_2^1}, \quad b = \frac{1}{1 + \gamma_0 \gamma_1^{25} \gamma_2^0},$$
$$c = -\frac{\gamma_0 \gamma_1^{40} \gamma_2^1}{1 + \gamma_0 \gamma_1^{40} \gamma_2^1}, \quad d = -\frac{\gamma_0 \gamma_1^{25} \gamma_2^0}{1 + \gamma_0 \gamma_1^{25} \gamma_2^0}.$$

Row-reduction leads to two equations in four variables:

$$a + c = 0 \quad \text{and} \quad b + d = 0$$

There are *infinitely many real solutions* to this system:

$$\gamma_0 = \gamma_1^{-25}, \quad \gamma_2 = \gamma_1^{-15}, \quad \gamma_1 > 0$$

This is not surprising, for we were given uninformative data:

$$\begin{array}{cccc} 40 & 25 & 40 & 25 \\ 1 & 0 & 1 & 0 \end{array}$$

A rigorous estimation method should not be able to provide unique estimates from such data.

Is it possible to maximize $L(\gamma_1^{-25}, \gamma_1, \gamma_1^{-15})$ w.r.t. γ_1 and describe the root surface corresponding to each γ_1 ?

If we were given the data on individuals 16-20 and 31-35 only, then the likelihood equations are

$$\begin{pmatrix} 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 & 1 \\ 21 & 46 & 32 & 23 & 25 & 23 & 47 & 57 & 25 & 60 \\ 0 & 1 & 0 & 1 & 0 & 1 & 0 & 1 & 1 & 1 \end{pmatrix} \begin{pmatrix} a_1 \\ \vdots \\ a_{10} \end{pmatrix} = 0$$

where

$$a_1 = \frac{1}{1 + \gamma_0 \gamma_1^{21} \gamma_2^0}, \dots, a_{10} = -\frac{\gamma_0 \gamma_1^{60} \gamma_2^1}{1 + \gamma_0 \gamma_1^{60} \gamma_2^1}$$

Load the data into *Macaulay2*, a software package for numerical algebraic geometry

Let a laptop computer run for hours

Macaulay2 finds *all* 1,346 complex solutions

Only 3 of the 1,346 solutions are real

Only 1 of the 3 real solutions has all components positive:

$(87982.8, 0.751485, 0.0197566)$

Conclusion: $(87982.8, 0.751485, 0.0197566)$ is the unique MLE.

Macaulay2 has therefore proved that the MLE exists and is unique.

The General Case

Suppose that the x_{ij} are integers (e.g., the Donner data) or rational numbers.

The ML equations reduce to a system of polynomial equations.

The Fundamental Theorem of Algebra: Every non-zero, one-variable polynomial of degree n , with complex coefficients, has exactly n complex roots (counted with multiplicity).

Rothe (1608), Euler (1749), Lagrange (1772), Laplace (1795), Gauss (1799), Argand (1806), Ostrowski (1920), ...

How does the Fundamental Theorem of Algebra generalize to several variables?

1841: F. Minding generalizes the FTA to two variables.

1975: D. Bernstein generalizes the FTA to arbitrary number of variables.

Bernstein's proof motivated numerical algorithms for sweeping through the values of the polynomial system to find *all* complex *isolated* roots.

Polynomial Homotopy Continuation algorithms

J. Verschelde, Univ. Illinois at Chicago: Extensive PHC website with software, examples, manuals, free downloads.

Garcia-Puente, Gross, Kahle, Petrović, Stasi, Sommese: People who know how to apply the software

Buot and Richards (2006). Counting and locating the solutions of polynomial systems of maximum likelihood equations, I. *J. Symbolic Computation*

Buot, Hoşten, and Richards (2007). Counting and locating . . . , II: The Behrens-Fisher problem. *Statistica Sinica*

Cox, Little, and O'Shea (1998). *Using Algebraic Geometry*, Springer

Gross, Drton, and Petrović (2012). The maximum likelihood degree of variance component models. *Electron. J. Statist.*

Sturmfels (1998). Polynomial equations and convex polytopes. *Amer. Math. Monthly*

As $n \rightarrow \infty$, the number of roots of ML equations does not always converge to 1

Problem: Estimate the *correlation* matrix of a multivariate normal distribution

Social scientists wish to estimate tetrachoric and polychoric correlations.

Constrained estimation problems; more difficult than estimating the covariance matrix.

This problem cannot be solved by estimating each bivariate correlation separately.

We must parametrize the set of correlation matrices carefully.

$N_3(0, R)$, a trivariate normal distribution with mean 0 and correlation matrix R

Collect a random sample and write down the likelihood function.

We solve the likelihood equations using *Bertini*, a software package for numerical algebraic geometry.

The likelihood equations seem to *always* have 35 complex solutions.

The number of statistically relevant solutions varies from 5 to 9.

Even with $n = 10^7$, we found cases with 9 statistically relevant solutions.

Conclusions

Statisticians often have complicated estimating equations with:

- Small sample sizes

- Large numbers of parameters

- Multiple roots

We recommend the use of *numerical algebraic geometry*

- 21st-century mathematical methods

- Powerful algorithms for solving estimating equations

- These algorithms compute *all* solutions of the equations